

Geographic Variation of Adjuvant Breast Cancer Therapy Initiation in the United States: Lessons From Medicare Part D

John A. Charlson, MD^{a,b}; Emily L. McGinley, MS, MPH^b; Ann B. Nattinger, MD, MPH^{a,b}; Joan M. Neuner, MD, MPH^{a,b}; and Liliana E. Pezzin, PhD, JD^{a,b}

Abstract

Background: Drug utilization under Medicare Part D varies significantly by geographic region. This study examined the extent to which geographic variation in Part D plan characteristics contributes to the variation in choice of initial endocrine therapy agent among women with incident breast cancer. **Methods:** Two-stage multivariate regression analyses were applied to the 16,541 women identified from Medicare claims as having incident breast cancer in 2006–2007. The first stage determined the effect of state of residence on the probability of having an aromatase inhibitor (AI), as opposed to tamoxifen, as initial endocrine therapy. The second stage provided estimates of the impact of state-specific Part D plan characteristics on variation in choice of initial therapy. **Results:** There was substantial residual geographic variation in the likelihood of using an AI as initial endocrine therapy, despite controlling for socioeconomic status, breast cancer treatment, and other factors. Regression-adjusted probabilities of starting an AI ranged from 57.3% in Wyoming to 92.6% in the District of Columbia. Results from the second stage revealed that variation in characteristics of Part D plans across states explained approximately one-third (30%) of the state-level variability in endocrine therapy. A higher number of plans with cost-sharing above the mean, greater spread in deductibles, and a greater spread in monthly drug premiums were associated with lower adjusted state probabilities of initiating an AI. In contrast, a higher number of drug plans with monthly premiums above the state mean and higher mean cost-sharing (in dollars) were both positively associated with likelihood of starting on an AI. **Conclusions:** Study findings suggest that variation in benefit design of Part D plans accounts for an important share of the large and persisting variability in use of AIs—the preferred oral therapy for breast cancer.

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The Medicare Prescription Drug, Improvement, and Modernization Act of 2003 authorized optional prescription drug coverage (Part D) for eligible beneficiaries beginning January 1, 2006. Unlike Medicare Parts A and B, Medicare coverage of prescription drugs relies on private plans within the various states; these plans are allowed wide discretion when setting features and prices. In order to ensure some degree of consistency across states, the Part D legislation included substantial regulatory oversight requiring plans to offer actuarially equivalent benefits, provide coverage of certain medi-

cations, and supply coverage information for a single Centers for Medicare & Medicaid Services (CMS) Web site to facilitate consumer comparison. A number of recent studies, however, suggest that drug utilization under Part D has been uneven, with prescription drug use among Part D enrollees varying significantly by geographic regions.^{1,2}

Adjuvant endocrine therapy for breast cancer is a useful model for studying the relationship between pharmaceutical plan characteristics and choice of therapy. For most postmenopausal women with hormone

From the ^aDepartment of Medicine and ^bCenter for Patient Care and Outcomes Research, Medical College of Wisconsin, Milwaukee, Wisconsin. Submitted April 7, 2016; accepted for publication August 2, 2017. The authors have disclosed that they have no financial interests, arrangements, affiliations, or commercial interests with the manufacturers of any products discussed in this article or their competitors. **Author contributions:** Study concept and design: Charlson, Pezzin,

Nattinger, McGinley, Neuner. *Data acquisition:* Pezzin, Nattinger. *Data analysis and interpretation:* Charlson, Pezzin, Nattinger, McGinley, Neuner. *Manuscript preparation:* Charlson, Pezzin, Nattinger. *Critical revision:* Pezzin, Neuner, McGinley, Charlson. Correspondence: John A. Charlson, MD, Center for Patient Care and Outcomes Research, Medical College of Wisconsin, 8701 Watertown Plank Road, Milwaukee, WI 53226. E-mail: jcharlso@mcw.edu

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receptor–positive (HR+) disease there are essentially 2 options for oral endocrine therapy: tamoxifen, whose efficacy was initially demonstrated in the mid-1980s,³ or an aromatase inhibitor (AI). Beginning in 2005, recommendations from ASCO suggested that adjuvant therapy for postmenopausal women with HR+ breast cancer include an AI, either alone or in sequence after tamoxifen.⁴ For much of the past decade, AI agents have been significantly more expensive than tamoxifen, with average costs exceeding \$300 USD monthly.⁵ Although AI agents recently became available in generic formulations with a considerably lower cost to the patient, historical patterns of use of breast cancer–specific endocrine therapies provide useful insights into the ways patients respond to various plan features. In addition, given the expanding number of expensive oral chemotherapy agents and the fact that several malignancies, such as chronic myeloid leukemia (CML) and renal cell carcinoma, have multiple oral treatment options available, our study provides an analytical framework that is applicable to other situations.

This article examines the extent to which geographic variation in Medicare Part D plan characteristics contribute to the variation in choice of initial endocrine therapy agent among women with early-stage breast cancer enrolled in Part D. To that end, we first quantified the *net* contribution of state of residence in women's probability of choosing the more expensive AIs, as opposed to tamoxifen, as the initial hormonal therapy agent. In a second-stage estimation, we examined whether state pharmaceutical plans' characteristics, such as annual AI drug costs, deductibles, premiums, and cost-sharing, significantly explain such state effects.

Methods

Sample Selection

The main sources of data for cohort selection were derived from Medicare medical and pharmaceutical claims. Our inclusion criteria consisted of women aged 65 to 89 years old identified from Medicare claims as having undergone incident breast cancer surgery in 2006 or 2007, based on a validated algorithm⁶ applied to nationwide Medicare Parts A and B claims. Subjects were required to be enrolled in Medicare for at least 12 months before the incident breast cancer surgery date in order to enable the

measurement of comorbidities. Subjects were further required to have an identified surgeon for their surgery, to have been enrolled in Medicare Part D for 12 months after the date of surgery, and to have started hormonal therapy with an AI (anastrozole, exemestane, letrozole) or tamoxifen during that period (following recommendations from the National Quality Forum and ASCO), leaving a sample of 16,541 women.

Variable Definitions

Initial (within 12 months from surgery) endocrine therapy drug choice (AI vs tamoxifen) was identified from individual-level pharmacy event records in the Medicare Part D data set. Information on age, race,⁷ poverty status (defined by enrollment in the Medicaid program or receipt of the federal low-income subsidy [LIS] benefit), and state of residence at the time of incident breast cancer surgery were derived from Medicare enrollment files. The number of comorbid conditions was characterized by examining individual inpatient, outpatient, and hospital provider claims for the 12-month period preceding the breast cancer surgery using the methodology described by Klabunde et al.⁸ Information about initial breast cancer treatment (mastectomy or breast-conserving surgery) and receipt of adjuvant chemotherapy was also obtained from Medicare claims.

Patients' sociodemographic and economic status was further characterized by including a zip code–level measure of per capita income obtained from the US Census Bureau,⁹ a variable capturing the density of women aged ≥ 65 years in the county obtained from the Area Health Resource File (AHRF),¹⁰ and residence in a rural area, defined according to the Core-Based Statistical Area definitions.^{10,11} County-level measures of density of radiation oncologists per elderly persons, obtained from the AHRF,¹⁰ were also considered for inclusion in the analyses as a proxy for access to oncology specialists, a factor potentially associated with differential patterns of care.

To characterize drug plan variation, we assembled information on the Part D Prescription Drug Plan (PDP) available to Medicare beneficiaries in each state and the District of Columbia by manually querying the CMS Web site used by Medicare beneficiaries to find and compare the Part D PDPs available in their areas.¹² For each plan in each state, information was collected on deductible amounts,

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monthly drug premiums, dollar range of drug cost sharing, and annual drug cost (ie, total out-of-pocket cost to a beneficiary who took only that particular endocrine agent at the recommended dosing for the entire year, excluding plan premiums) for each endocrine agent of interest (anastrozole, exemestane, letrozole, and tamoxifen).⁵ Based on these data, we assessed variability among plans at the state level by calculating 4 measures—mean cost, spread between the top and bottom quartile costs, number of plans with cost below the state mean, and number of plans with cost above the state mean—for each of 5 constructs: (1) monthly drug premium, (2) deductible, (3) upper value of the dollar cost-sharing range, (4) lower value of the dollar cost-sharing range, and (5) cost after deductible for AIs. Finally, we calculated the total number of Part D PDPs and total PDP enrollment in each state. All values were determined for each of the 50 states (except for the District of Columbia).

Statistical Analysis

Basic descriptive statistics were used to characterize the sample population according to the subjects' sociodemographic and economic profile, health status, and initial endocrine therapy.

Our main analyses relied on multivariate regression models to test the hypothesis of a significant contribution of Part D plan characteristics to residual geographic variation in initial therapy modality. The multivariate analysis proceeded in 2 stages. In the first stage, a multivariate probit regression was used to examine the effect of state of residence on the probability of initiating AI therapy, controlling for the various patient factors described earlier.¹³ In this stage, the goal was to adjust for all available individual and small-level covariates to allow the most conservative estimate possible of residual variation by state for the following stage. Robust standard errors were computed to account for clustering (ie, multiple observations on patients treated by the same surgeon) and heteroskedasticity effects.

The state “fixed effects,” captured by each state's estimated coefficient in this first-stage individual-level estimation, were then used as the dependent variable in the second stage estimation examining the effect of state-level Part D plan characteristics, as well as number of and enrollment in Part D PDPs, on state variation in AI use as the initial endocrine

therapy. Given that we had no a priori hypotheses regarding which specific Part D plan characteristics might explain geographic variation in AI use, we relied on a forward stepwise selection, and the magnitude/significance of the Wald likelihood ratio test statistic of alternative specifications, to determine the best fit/most explanatory model.

A *P* value of <.05 was considered the critical level to determine statistical significance in the first-stage regression. However, a lower threshold (*P*<.15) was used in the second-stage fixed-effects regression, wherein the sample size was reduced to 50 states. Associations approaching significance can provide further insight into Part D plan characteristics effects on state variation in therapy initiation drug for which the conventional *P*<.05 level is not realized because of small sample size. All analyses, which were approved by the relevant Institutional Review Boards, were conducted using SAS version 9.3 (SAS Institute Inc., Cary, NC) and Stata version 12.0 (StataCorp LP, College Station, TX).

Results

Sample Characteristics

Of the cohort of 16,541 breast cancer survivors who received endocrine therapy in the first year after incident breast cancer surgery, nearly half were aged <75 years, most were Caucasian, and 26.8% had annual incomes sufficiently low to warrant receipt of Medicare D LIS. The cohort was generally healthy, with more than half having no comorbidities. Most received breast-conserving surgery, and 16.7% were also treated with adjuvant chemotherapy; most used an AI as their initial endocrine therapy (Table 1).

However, there was substantial variation by state, with the unadjusted state proportion of women initiating hormonal therapy with an AI ranging from 54.1% in Wyoming to 83.3% in the District of Columbia (unadjusted results not shown).

Geographic Variation and Other Factors Associated With Initial AI Use

As shown in Table 1, there were a number of differences between those who initiated an AI and those who began tamoxifen. Several of the bivariate differences across the 2 groups persisted in the multivariate regression (Table 2). Women who lived in neighborhoods ranking in the top quartile of per capita

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Table 1. Summary Statistics

	Overall, % (n=16,541)	Initial Endocrine Therapy Drug	
		AI, ^f % (n=12,654; 76.5%)	Tamoxifen, % (n=3,887; 23.5%)
Age, y ^a			
65–74	48.9	49.2	48.2
75–84	42.2	42.3	41.9
85–89	8.8	8.5	9.8
Race ^a			
African American	6.4	6.5	5.9
Hispanic	3.9	4.1	3.2
Non-African American, non-Hispanic	89.7	89.4	90.9
Comorbidities ^a			
0	55.8	55.2	57.8
1	26.8	27.1	25.9
≥2	17.4	17.7	16.3
Type of breast surgery ^a			
Mastectomy	40.5	42.4	34.3
Breast-conserving surgery	59.5	57.6	65.7
Surgery year			
2006	37.8	38.0	37.2
2007	62.2	62.0	62.8
Adjuvant chemotherapy ^a	16.7	19.3	8.5
Poverty status			
Low-income subsidy recipient ^a	26.8	27.9	23.2
Per capita income in zip code ^a			
Poorest quartile	24.1	24.0	24.2
Middle quartiles	50.3	49.3	53.4
Wealthiest quartile	25.7	26.7	22.4
Residence in rural county ^a	27.8	26.5	32.1
Density of elderly women in county, mean (SD) ^a	0.80 (0.23)	0.79 (0.22)	0.82 (0.23)

Abbreviation: AI, aromatase inhibitor.

^aStatistical significance at the $P < .05$ level based on Pearson chi-square test statistics.

income and those who had a LIS were significantly more likely to start on an AI, as were women who received chemotherapy and those who underwent a mastectomy. Women living in rural areas and those residing in a county with a greater density of elderly women were less likely to initiate AI therapy, whereas those residing in a county with a greater density of radiation oncologists were more likely to initiate AI therapy.

Parameter estimates for the first stage estimation were used to compute the adjusted probabilities that a woman living in a given state would start an AI as her initial endocrine therapy, controlling for all other factors. After controlling for clustering and po-

tentially confounding variables, such as differences in the subject composition or rurality of states, the adjusted probabilities ranged from a low of 57.3% in Wyoming to a high of 92.6% in the District of Columbia (Figure 1). North Carolina was selected as the reference state, because it most closely approximated the unadjusted national mean (70.6%) AI use for the sample. Using a >5% deviation from the national mean in either direction as threshold, the other states were determined to have AI initiation rates either significantly above (colored in blue), significantly below (colored in red), or not different (within a 5% margin, colored in grey) from the reference state of North Carolina. These differences represent

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the *independent* effect of state of residence on choice of initial endocrine therapy, after controlling for initial breast cancer treatment, individual and ecologic measures of socioeconomic status (SES), and the other potential cofounders. A total of 22 states had AI initiation rates significantly higher than the reference state (national mean), whereas 12 states, mostly in the Northwest and Midwest regions, had AI initiation rates significantly below.

State Variation in Initial Endocrine Therapy and Part D Plan Characteristics

To determine the effect of state-level Part D plan characteristics on state variation in AI use, a second stage estimation was conducted, using the state “fixed effects” from the first stage analysis as the dependent variable. Table 3 shows the results from the best-fitting stepwise linear regression. Variation in characteristics of Part D plans across states explained 30% of the overall state-level variability in initiating hormonal therapy with an AI versus tamoxifen. States with a higher number of plans with monthly premiums above the state mean had higher rates of AI initiation, whereas states with a higher number of plans with cost-sharing above the mean had lower rates of AI initiation. States with higher mean cost-sharing (in USD) were also associated with higher probabilities of AI initiation. Finally, a greater spread in deductible or monthly drug premium costs was associated with lower adjusted probabilities of AI as initial therapy among women in those states.

Discussion

Substantial residual geographic variation was observed in the likelihood of using an AI as initial endocrine therapy among this nationally representative cohort of 16,541 older breast cancer survivors, despite controlling for patient SES, breast cancer treatment, and other potential confounders. The adjusted probability of choosing an AI versus tamoxifen as initial adjuvant endocrine therapy ranged from 57.3% in Wyoming to 92.6% in the District of Columbia. Nearly one-third of such (adjusted) state-level variation was explained by differences in characteristics of Medicare Part D plans available in the subjects’ state of residence.

Our results indicate that a substantial portion (30%) of state-to-state variation in the probability

Table 2. Multivariate Probit Results for First-Stage, Patient-Level Estimation^a

Variable	Coefficient (P Value)
Age, y	
65–74	Ref
75–84	0.030 (.207)
85–89	–0.058 (.144)
Race	
Non-African American, non-Hispanic	Ref
African American	–0.042 (.398)
Hispanic	–0.052 (.400)
Comorbidity level	
0	Ref
1	0.028 (.294)
≥2	0.040 (.200)
Low-income subsidy beneficiary	
No	Ref
Yes	0.119 (.000)
Residence in high per capita income zip code	
No	Ref
Yes	0.091 (.002)
Residence in rural county	
No	Ref
Yes	–0.035 (.307)
Surgery type	
Breast-conserving surgery	Ref
Mastectomy	0.183 (.000)
Surgery year	
2006	Ref
2007	–0.027 (.228)
Chemotherapy	
No	Ref
Yes	0.507 (.000)
Density of radiation oncologists in county	0.238 (.041)
Density of elderly women in county	–0.179 (.010)
Wald likelihood ratio test ^b	574 (.000)

^aThe estimation also adjusted for the patient’s state of residence (coefficients not shown; see Figure 1 for state-specific adjusted probabilities of having an aromatase inhibitor [AI] as the initial endocrine therapy drug of choice). Positive (negative) coefficients indicate factors that increase (decrease) the likelihood that a patient with that trait would start endocrine therapy with an AI as opposed to tamoxifen. Standard errors have been adjusted to account for clustering effects (ie, multiple observations on patients for a given surgeon).

^bThe Wald likelihood ratio test, a statistic commonly used for testing the joint significance of the explanatory variables, was used to compare and inform model selection.

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Table 3. Effect of Characteristics of Part D Plans on State Variation

State-Specific Part D Plans Characteristic	Coefficient (SE)	P Value
Number of drug plans with monthly premium above the mean	0.075 (0.020)	<.0001
Mean drug cost-sharing (top quartile value, in dollars)	0.091 (0.030)	.003
Spread between the top and bottom quartiles of deductible	-0.021 (0.007)	.007
Number of drug plans with cost-sharing above the mean	-0.069 (0.026)	.01
Spread between the top and bottom quartiles of the monthly drug premium	-0.023 (0.012)	.07
Adjusted R ²	0.30	

The dependent variable in this second-stage estimation is the state fixed effects obtained from the first-stage individual-level estimation (n=50). These fixed effects capture the independent influence of state of residence on the probability of choosing an aromatase inhibitor as initial endocrine therapy, after controlling for potential patient-level confounders and clustering effects. Results are based on forward stepwise regression, applied to a larger pool of state-level variables, which included all variables capturing variation in benefit plan design (see text for detailed description), as well as number of Part D Prescription Drug Plans (PDPs) in each state and PDP enrollment in each state. None of these additional variables were selected for inclusion for the best-fitting stepwise likelihood maximization process.

idence. It is likely that a portion of the residual unexplained variation in medication utilization is due to practice style of prescribing providers who are correlated with geographic location.^{22,23} For example, AI use in our cohort is correlated with a more aggressive approach to treatment, as evidenced by the positive associations between initial AI choice, chemotherapy, and mastectomies. There is no reason to believe that factors such as tumor size, stage, HER2 status, and HR status, which drive these choices in treatment, would vary so widely from state to state. In addition, even after controlling for urban setting and population of elderly women, we observed a higher utilization of AI therapy in areas with a greater density of cancer specialists (radiation oncologists). It is also possible that direct and indirect drug marketing efforts impact the variability we observed.

ASCO has recommended since 2005 that adjuvant treatment for postmenopausal women with HR+ breast cancer should include an AI.⁴ At the time our cohort of patients was starting treatment, there were large high-quality studies^{24,25} that showed longer disease-free survival and fewer distant recurrences in women treated for 5 years with AIs compared with tamoxifen. Other studies published around the same time demonstrated that switching to an AI after 2 to 3 years of tamoxifen was superior to taking tamoxifen alone for 5 years, but studies documenting that upfront AI for 5 years and sequential tamoxifen and AI had similar outcomes were not published until after 2010. Therefore, in 2006 and 2007, physicians were left to “guess” which strategy was better for patients, perhaps using their assessments of the importance of the additional benefit of AIs, such as reduction in several commonly recognized adverse effects associ-

ated with tamoxifen, including thromboembolism and endometrial carcinoma.²⁶ Data from our study indicate that 76.6% of patients started on an AI as initial endocrine therapy, suggesting rapid adoption of this class of medications as preferred initial therapy. The NCI now states in its Health Professional guidelines that AIs have become the first-line adjuvant therapy for postmenopausal women.²⁷ Consistent with these statements, ASCO recommends use of tamoxifen alone only in instances where women have a contraindication to or are intolerant of AIs.⁴ There is no reason to think that contraindications to or intolerances of AIs would vary substantially at the state level.

Although evidence suggests that both the utilization of and adherence to prescription medications overall have increased since the implementation of the Medicare Part D program in 2006,^{28–30} the proportion of women prescribed AIs as initial treatment in this Part D era was remarkably similar to that reported for patients treated before the advent of the program.^{13,17,31} Our study finds that differences in treatment by age and SES persisted after implementation of Part D, with older patients and those with multiple comorbidities being less likely to initiate therapy with AIs. Consistent with previous studies documenting significant urban/rural differences in breast cancer treatment, women living in rural areas were also less likely to use AIs as their initial therapy.^{6,13,17}

Our study has a number of limitations that merit comment. Our analysis focuses on choice of an AI as the *initial* therapy. There are, however, 2 strategies for AI use in women in with early-stage estrogen receptor–positive breast cancer: upfront AI treatment or sequencing to an AI after 2 to 3 years of

tamoxifen treatment. Given that our analysis focuses on the former approach, we are unable to make any statements about the “appropriateness” of treatment. There are also legitimate clinical reasons not to prescribe the more expensive but more effective oral AI as initial therapy, such as the increased risk of osteoporosis associated with these agents. Unfortunately, we were unable to assess osteoporosis or differences in accessibility to its treatments in our study. It is unlikely, however, that concerns about the incidence of this or other AI-related side effects would explain the systematic and substantial variability in prescription patterns by state of residence observed in our sample. In addition, given that this study is part of a larger project focusing on the early years of Medicare Part D implementation, our analysis examines the experience of patients with breast cancer during 2006 to 2008, a time that precedes the availability of generic alternatives to AIs (2010–2011). Finally, we purposely opted against a more direct, individual-level analysis of the relationship between AI initiation and characteristics of Part D plans chosen by women in our sample. Given that nationwide administrative data sets, such as Medicare, do not contain information that would have enabled us to account and correct for participation (in Part D) and self-selection (in specific Part D plans) biases, any results from such a more direct approach would have been spurious. By relying on a mixed, individual-ecologic, 2-stage approach, we provide unbiased estimates of the effect of (differences in characteristics of) Part D plans available in each state on state-level variation in AI use, net of individual socioeconomic, demographic, and disease confounders.

Conclusions

This study raises concerns that, although Medicare Part D has improved older patients’ access to prescription drugs, it did relatively little to reduce geographic variation in treatment patterns. Variations in the characteristics of Part D plans available in different states accounted for an important share of the remaining geographic variability, a finding of increasing relevance in an era when there is a rapidly expanding number of oral anticancer therapies. Although the cost of AIs has declined substantially since the recent advent of generic alternatives, our findings provide an important framework for understanding the effect of Part D plan variability on treatment of other malignancies using expensive oral chemotherapy agents. A more recent example might be CML, which has numerous costly oral drugs approved for use, and until one of them (imatinib) recently came off patent, none of the drugs had affordable generic versions. Insurance plans should be structured to encourage utilization of the agents with best evidence of efficacy and to facilitate access to these drugs for the patients who need them.

Our findings also reveal significant variations in healthcare provision not explained by individual sociodemographic and disease characteristics, or differences in state-level access to care or Part D plan characteristics. Studies that strive to understand the causes of this unexplained variation will be important as we work to maximize value in the US healthcare systems.

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